

Uncertainty Quantification of GEOS-5 L-Band Radiative Transfer Model Parameters using Bayesian Inference and SMOS Observations

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Abstract

Uncertainties in L-band (1.4 GHz) microwave radiative transfer modeling (RTM) affect the simulation of brightness temperatures (Tb) over land and the inversion of satellite-observed Tb into soil moisture retrievals. In particular, accurate estimates of the microwave soil roughness, vegetation optical depth and scattering albedo for large-scale applications are difficult to obtain from field studies and often lack an estimate of uncertainty. Here, a Markov Chain Monte Carlo (MCMC) simulation method is used to determine satellite-scale estimates of RTM parameters and their posterior uncertainty by minimizing the misfit between long-term averages and standard deviations of simulated and observed Tb at multiple incidence angles, at horizontal and vertical polarizations, and for morning and evening overpasses. Tb simulations are generated with the land model component of the Goddard Earth Observing System (version 5) and confronted with Tb observations from

the Soil Moisture Ocean Salinity (SMOS) satellite mission. The maximum a posteriori density (MAP) parameter values reduce the root-mean-square differences between observed and simulated long-term Tb averages and standard deviations to 3.4 K and 2.3 K, respectively. The relative uncertainty of the posterior RTM parameter estimates is typically less than 25% of the MAP parameter value, whereas it exceeds 100% for literature-based prior parameter estimates. It is also shown that the parameter values estimated through Particle Swarm Optimization are in close agreement with those obtained from MCMC simulation. The MCMC results for the RTM parameter values and the uncertainties presented herein are directly relevant to the need for accurate Tb modeling in global land data assimilation systems.

¹ *Keywords:*

² radiative transfer modeling, brightness temperature, Bayesian parameter
³ estimation, uncertainty, Markov Chain Monte Carlo simulation, SMOS

1. Introduction

Uncertainties in radiative transfer modeling (RTM) affect the simulation of brightness temperatures (Tb) over land and the inversion of satellite-observed Tb to soil moisture retrievals. Quantification of these uncertainties is crucial to producing, validating and using L-band (1.4 GHz) passive microwave data, such as those obtained from the Soil Moisture Ocean Salinity (SMOS, Kerr et al. (2010)) and future Soil Moisture Active Passive (SMAP, Entekhabi et al. (2010)) missions. Yet, it is not particularly clear which RTM formulation and parameter values to use for large-scale applications.

In the context of large-scale forward Tb simulation, several studies have analyzed the effect of different RTM formulations for the microwave roughness length, vegetation parameterization and soil dielectric model (Drusch et al., 2009; de Rosnay et al., 2009), the impact of parameter values (De Lannoy et al., 2013) and the sensitivity to dynamic land surface variables (Balsamo et al., 2006). Similarly, soil moisture retrievals based on Tb observations are affected by the RTM formulation and parameter values (Crosson et al., 2005; Panciera et al., 2009; Konings et al., 2011; Parinussa et al., 2011), as well as by the choice of background and auxiliary fields, such as soil temperature and vegetation characteristics (Kerr et al., 2012; O'Neill et al., 2012). Collectively, these studies suggest that RTMs exhibit significant uncertainty and that the impact of this uncertainty on large-scale Tb simulations and soil moisture retrievals remains unclear.

Estimating microwave RTM parameters and their uncertainty is a major challenge, especially at larger spatial scales. Field experiments have provided RTM parameter values (de Rosnay et al., 2006; Grant et al., 2007; Panciera

1 et al., 2009; Sabater et al., 2011), but mostly without an underlying estimate
2 of their uncertainty. De Lannoy et al. (2013) derived global-scale RTM pa-
3 rameter values and ad hoc uncertainty estimates using SMOS observations
4 and Particle Swarm Optimization (PSO, Kennedy and Eberhart (1995)).
5 PSO is specifically designed to find the optimal parameter values, without
6 recourse to estimating their underlying uncertainty.

7 In this paper, we apply a (Bayesian) Markov chain Monte Carlo (MCMC)
8 simulation method to estimate the posterior RTM parameter distribution.
9 The DiffeRential Evolution Adaptive Metropolis (DREAM_(ZS)) algorithm is
10 used with parallel direction and snooker sampling from past states (Vrugt
11 et al., 2008, 2009; Laloy and Vrugt, 2012). Bayesian approaches such as
12 DREAM_(ZS) have several advantages over optimization methods such as PSO.
13 The explicit treatment and analysis of uncertainty help to understand which
14 parts of the RTM are well resolved and which elements require further atten-
15 tion. Furthermore, a formal analysis of the residuals can be used to check the
16 validity of our assumptions about the probabilistic properties of the errors
17 and to discern whether reliable parameter values have been derived.

18 Implementation of the Bayesian paradigm coupled with MCMC simu-
19 lation comes at an increased computational expense. Full exploration of
20 the posterior distribution is very costly, and hence difficult to warrant in
21 global scale operational applications that rely on evolving modeling systems
22 in need of frequent re-calibrations. Yet, this paper will demonstrate that
23 the proposed approach provides important insights into the uncertainty of
24 large-scale RTM parameters, and can be used to benchmark the results from
25 optimization algorithms such as PSO.

1 The goals of the present paper are thus to infer large-scale RTM param-
2 eters and their posterior uncertainty using a Bayesian method, and to study
3 the associated simulated Tb uncertainty. This research complements and
4 advances the work of De Lannoy et al. (2013) where PSO was used to find
5 parameters without formal statistical estimates of uncertainty. We are using
6 the Goddard Earth Observing System, version 5 (GEOS-5) modeling frame-
7 work that will be used to generate the planned global SMAP Level 4 Surface
8 and Root Zone Soil Moisture (L4_SM) data product through assimilation of
9 SMAP Tb observations (Reichle et al., 2012). Here, the focus is on estimat-
10 ing time-invariant RTM parameters and their uncertainty by minimizing cli-
11 matological (long-term) differences between multi-angular, horizontally and
12 vertically polarized Tb for morning and evening overpasses from SMOS ob-
13 servations and GEOS-5 simulations, while explicitly treating and estimating
14 their respective uncertainties.

15 The time-invariant estimated parameters will be used in a data assimila-
16 tion system (outside the scope of this paper), where state variables such as
17 soil moisture and soil temperature will be updated in response to short-term
18 variations in the observed Tb. Residual long-term Tb errors, or biases, that
19 remain after the estimation of the RTM parameters will be addressed through
20 model refinement and within the data assimilation system. The (Bayesian)
21 uncertainties of the RTM parameters and the corresponding estimates of the
22 residual observation and model error presented here will help with the devel-
23 opment of such data assimilation systems. Furthermore, the uncertainties of
24 the RTM parameters reported here will also facilitate sensitivity analyses of
25 soil moisture retrieval algorithms (Parinussa et al., 2011).

1 To summarize, in this paper we apply MCMC simulation using multi-
 2 angular SMOS Tb observations to (i) determine if the maximum a posteriori
 3 density (MAP) parameter values derived from the posterior distribution sam-
 4 pled with DREAM_(ZS) can be approximated using PSO, (ii) obtain reliable
 5 estimates of parameter uncertainty, and (iii) quantify the impact of errors in
 6 parameters and other sources on Tb simulations. The remainder of this paper
 7 is organized as follows. Section 2 summarizes the modeling system and the
 8 SMOS observations used in the present study. This is followed in section 3
 9 by a brief description of the DREAM_(ZS) MCMC simulation method. This
 10 section also discusses several quantitative diagnostic metrics to analyze the
 11 simulated Tb uncertainty. Section 4 discusses the main findings and results
 12 of this paper. This is followed in section 5 with a summary and conclusions.

13 **2. Observations and Model**

14 *2.1. SMOS Tb Data*

15 Since its launch in November 2009, the SMOS mission provides global
 16 L-band Tb data at a nominal spatial resolution of 43 km. On average, a
 17 given location on the equator is revisited once every 3 days. Here we use the
 18 multi-angular, full polarization Tb data from the period 1 July 2010 to 1 July
 19 2012. Specifically, the data are extracted from the MIR_SCLF1C product,
 20 with processor version 504 for the years 2010 and 2011 (reprocessed in 2012),
 21 and version 551 from January 2012 onwards. De Lannoy et al. (2013) discuss
 22 in detail the various steps involved in the processing of the SMOS data. Most
 23 importantly, the data are screened extensively using both product-based data
 24 quality information and model-based quality control rules. Furthermore, the

1 data are spatially mapped onto a 36 km Equal-Area Scalable Earth (EASE)
2 grid and binned per incidence angle. Consistent with De Lannoy et al. (2013),
3 only a subset of 6 incidence angles is used: $\theta=[32.5^\circ, 37.5^\circ, 42.5^\circ, 47.5^\circ, 52.5^\circ,$
4 $\text{and } 57.5^\circ]$, where, for example, 32.5° represents the average of all Tb data
5 with incidence angles between 32 and 33 degrees.

6 To estimate the microwave RTM parameters, long-term averages (\mathbf{m}_o)
7 and standard deviations (\mathbf{s}_o) of the SMOS data are computed separately for
8 each of the 6 incidence angles, 2 polarizations (horizontal H and vertical V),
9 and 2 overpass times (ascending at $\sim 06:00$ h local time (LT), descending at
10 $\sim 18:00$ h LT). This results in a total of 48 “observations” per grid cell: 24
11 ($= 6 \times 2 \times 2$) observations of the long-term average Tb and 24 observations
12 of the long-term Tb standard deviation. Section 3 provides a more extensive
13 description of how these 48 observations are used.

14 2.2. GEOS-5 Tb Modeling

15 The modeling combines (i) land surface modeling with the Catchment
16 land surface model (CLSM) and (ii) radiative transfer modeling with a tau-
17 omega model to simulate long-term Tb averages and standard deviations. As
18 in De Lannoy et al. (2013), the GEOS-5 CLSM (Koster et al., 2000) is set up
19 on the 36 km EASE grid and spun up prior to the SMOS observation period.
20 Surface meteorological forcing data at a $1/2^\circ \times 2/3^\circ$ spatial and hourly tem-
21 poral resolution are taken from the Modern-Era Retrospective analysis for
22 Research and Applications (MERRA, Rienecker et al. (2011)). The MERRA-
23 precipitation is corrected with gauge-based precipitation from the National
24 Oceanic and Atmospheric Administration (NOAA) Climate Prediction Cen-
25 ter “Unified” (CPCU) product (Reichle, 2012). The model version is the

1 same as that used for the MERRA-Land data product (Reichle et al., 2011),
 2 except for two changes that more closely align the model with the version
 3 that will ultimately be used for the SMAP L4_SM data product: (i) the sur-
 4 face soil moisture pertains to the top 5 cm surface layer (as opposed to the
 5 top 2 cm layer in MERRA-Land), and (ii) a preliminary version of updated
 6 soil parameters from a forthcoming version of GEOS-5 is used.

7 The vegetation parameterization in CLSM uses 8 default vegetation classes.
 8 For the RTM simulations, these classes are further refined into the 16 classes
 9 defined by the Moderate Resolution Imaging Spectroradiometer (500 m MOD12Q1
 10 V004) International Geosphere-Biosphere Programme (IGBP) land cover
 11 classification (Loveland and Belward, 1997). Figure 1 shows the North Amer-
 12 ican study domain which includes 9 of the 16 IGBP vegetation classes.

13 In essence, T_b is determined by the surface soil temperature, soil moisture
 14 and other soil and vegetation characteristics. To simulate L-band T_b , the
 15 prognostic CLSM soil moisture, soil temperature, vegetation water content,
 16 air temperature and climatological vegetation dynamics are used as inputs
 17 to a diagnostic zero-order (τ - ω) microwave RTM, briefly described
 18 in Appendix A. The key model parameters that determine the microwave
 19 surface roughness h , the scattering albedo ω , and vegetation optical depth
 20 τ will be estimated using the multi-angular SMOS observations (section 3).
 21 As outlined in Appendix A, h is a function of soil moisture and the time-
 22 invariant parameters h_{min} and h_{max} (Eq. A.4), and τ depends on the leaf
 23 area index (LAI) and the time-invariant parameters b_H and b_V (Eq. A.6).

1 **3. Methods**

2 *3.1. Overview*

3 Keeping up with De Lannoy et al. (2013), the objective of the param-
4 eter estimation is to minimize the differences in long-term (climatological)
5 averages and standard deviations between multiple types of SMOS-observed
6 and GEOS-5-modeled Tb. We purposely do *not* minimize differences in the
7 time domain as the goal of the present paper is to derive parameter estimates
8 that result in the smallest possible long-term bias in the simulation of Tb.
9 Short-term differences between Tb observations and simulations will be ad-
10 dressed in future studies using sequential data assimilation. We estimate a
11 time-invariant multi-dimensional parameter set, hereafter referred to as α ,
12 that determines climatological features of the simulated Tb. The parame-
13 ters are estimated locally, i.e., for each grid cell independently, and only for
14 non-frozen land surface conditions as determined by the GEOS-5 modeling
15 system.

16 In addition to finding parameter estimates that result in accurate Tb sim-
17 ulations, the Bayesian methodology used here samples the complete posterior
18 parameter distribution. This distribution summarizes the dispersion of the
19 parameters, and can be used to derive the model simulation uncertainty by
20 propagating each sample of the posterior through the model. If our proba-
21 bilistic assumptions about the constituent error sources that affect the Tb
22 simulations are correct, then our analysis will provide maximum a posteriori
23 density (MAP) parameter estimates that accurately reflect their true values,
24 and the underlying posterior uncertainty will be statistically meaningful. In
25 other words, we seek accurate and precise parameter values with a minimum

1 simulation bias, and statistically meaningful spread.

2 Table 1 gives an overview of the parameters estimated in different ex-
3 periment cases. All cases estimate the 5 most relevant RTM parameters:
4 h_{min} , $\Delta h \equiv h_{max} - h_{min}$, b_H , $\Delta b \equiv b_V - b_H$ and ω . Simultaneous inference
5 of these select parameters accounts for their correlation in minimizing the
6 error residuals, and such an approach is thus preferred over sequential (step-
7 wise) fitting of the individual parameters (De Lannoy et al., 2013). Based
8 on these *time-invariant* parameters, *time-variant* values of h , τ_H and τ_V are
9 computed, using dynamic information about soil moisture for h (Eq. A.4)
10 and LAI for τ (Eq. A.6). Time-averaged results for $\langle h \rangle$ and $\langle \tau \rangle$ are
11 then presented, where $\langle \cdot \rangle$ denotes the long-term time average. The RTM
12 parameters are estimated with DREAM_(ZS) (section 3.2) or PSO (Appendix
13 B), hereafter referred to as cases D and P, respectively, both of which esti-
14 mate h_{min} , Δh , b_H , Δb and ω . Furthermore, a third case, D _{σ} , additionally
15 estimates the residual Tb error statistics σ_m and σ_s , using DREAM_(ZS) (dis-
16 cussed below). We thus estimate 5 parameters per grid cell for cases P and
17 D, and 7 parameters per grid cell for case D _{σ} .

18 To derive these parameters, we minimize per grid cell the *climatological*,
19 or long-term, differences between 48 Tb observations and simulations. These
20 $2 \times 24 = 48$ observations consist of long-term Tb averages and Tb standard
21 deviations for the 24 combinations of 2 polarizations, 2 overpass times, and 6
22 incidence angles. For simplicity, the errors in these observations are assumed
23 to be independent, that is, we neglect correlations in instrument errors and
24 errors between H- and V-polarized observations at identical incidence angles.
25 Similarly, the simulation errors are assumed to be independent, even though

1 some correlation is to be expected. Cross-correlations in observation or simu-
2 lation errors between various incidence angles would increase the uncertainty
3 in the posterior parameter estimates. Note that temporal correlations in the
4 errors are of little concern because the observations are long-term averages
5 and standard deviations, and not measurements in the time domain (Wöhling
6 and Vrugt, 2011).

7 In keeping up with De Lannoy et al. (2013), the two years of historical
8 SMOS data are divided into a calibration period (1 July 2011 - 1 July 2012)
9 and an evaluation period (1 July 2010 - 1 July 2011). To ensure a meaningful
10 calibration at each grid cell, we impose a minimum of 20 valid data points
11 (N_i) per year and per overpass time to compute the long-term Tb average and
12 standard deviation for a particular combination ($i = 1, \dots, 24$) of incidence
13 angle, polarization and overpass time. The requirement of $N_i \geq 20$ is used for
14 the calculation of evaluation statistics as well. In addition, for the calibration,
15 we always require a minimum total number of data points of $480=20 \times 6$
16 angles \times 2 polarizations \times 2 overpass times.

17 This study relies on proper cross-validation using independent calibration
18 and evaluation periods. Preparatory work for the forthcoming operational
19 SMAP L4_SM product uses all available historical SMOS data in the cali-
20 bration, and the results are comparable to what is reported herein. Given
21 the similar nature of L-band observations from SMOS and SMAP, we expect
22 that parameters calibrated with SMOS observations will initially serve well
23 in the SMAP L4_SM system, while a sufficient SMAP data record is being
24 accumulated for later re-calibration.

1 3.2. Markov Chain Monte Carlo (MCMC) Sampling

2 The Bayesian paradigm provides a framework for the treatment of all
 3 sources of uncertainty in modeling Tb. In this paper, we focus on parameter
 4 uncertainty, and treat the other sources of error as a single lumped term
 5 (details to follow). The posterior probability distribution of the parameters
 6 is computed by combining the observation likelihood $p(\mathbf{m}_o, \mathbf{s}_o|\boldsymbol{\alpha})$ with a prior
 7 distribution $p(\boldsymbol{\alpha})$:

$$p(\boldsymbol{\alpha}|\mathbf{m}_o, \mathbf{s}_o) = \frac{p(\mathbf{m}_o, \mathbf{s}_o|\boldsymbol{\alpha})p(\boldsymbol{\alpha})}{\int_{\boldsymbol{\alpha}} p(\mathbf{m}_o, \mathbf{s}_o|\boldsymbol{\alpha})d\boldsymbol{\alpha}} \quad (1)$$

8 The observations consist of long-term averages ($m_{i,o} \in \mathbf{m}_o$) and standard
 9 deviations ($s_{i,o} \in \mathbf{s}_o$) of Tb for 24 different combinations of incidence angles,
 10 polarizations and overpass times ($i = 1, \dots, 24$). The denominator is a nor-
 11 malization factor and thus it suffices to maximize $p(\mathbf{m}_o, \mathbf{s}_o|\boldsymbol{\alpha})p(\boldsymbol{\alpha})$ to find
 12 the posterior distribution of $\boldsymbol{\alpha}$. In practice, it is difficult to solve this prob-
 13 lem analytically and we therefore resort to MCMC simulation to generate a
 14 sample of the posterior target distribution.

15 In this paper, the DREAM_(ZS) algorithm (Vrugt et al., 2008; Laloy and
 16 Vrugt, 2012) with sampling from past states is used to efficiently explore
 17 the posterior parameter distribution. This algorithm adaptively updates the
 18 scale and orientation of the proposal distribution during sampling, and is
 19 specifically designed to rapidly explore multi-dimensional target distribu-
 20 tions. In DREAM_(ZS), multiple chains are running in parallel and the up-
 21 date of a chain is determined from an external sample of points that col-
 22 lectively summarizes the search history of all the individual chains. The
 23 log-likelihood of the current and proposed parameter values are compared

1 using the Metropolis selection rule. If the proposal is accepted, the chain
 2 moves to this new point, otherwise the chain remains at its current position.
 3 Diminishing adaptation of the external archive of samples ensures conver-
 4 gence to the exact posterior distribution.

5 We assume a Gaussian prior for each of the individual parameters $\alpha_{0,k} \in$
 6 $\boldsymbol{\alpha}_0$. The mean and standard deviation of this multi-normal distribution $p(\boldsymbol{\alpha})$
 7 are derived from literature values summarized in Table 1. Note that these
 8 values were referenced as ‘Lit2’ in De Lannoy et al. (2013). The prior RTM
 9 parameters were subjectively selected out of a range of possible values in
 10 the literature. The prior mean of parameter k is given by a (vegetation-
 11 dependent) value $\alpha_{0,k}$ and the standard deviation $\sigma_{\alpha_{0,k}}$ is computed as $\sigma_{\alpha_{0,k}}^2 =$
 12 $(\alpha_{max,k} - \alpha_{min,k})^2/12$, where $\alpha_{max,k}$ and $\alpha_{min,k}$ denote the respective upper
 13 and lower bounds.

14 The following log-likelihood function is used to minimize the differences
 15 in long-term Tb averages and standard deviations between the observations
 16 $(m_{i,o}, s_{i,o})$ and corresponding simulations $(m_i(\boldsymbol{\alpha}), s_i(\boldsymbol{\alpha}))$:

$$L = \ln(p(\mathbf{m}_o, \mathbf{s}_o | \boldsymbol{\alpha})) = -\frac{24}{2} \ln(2\pi) - \frac{1}{2} \sum_{i=1}^{24} \ln(\sigma_{i,m}^2) - \sum_{i=1}^{24} \frac{(m_{i,o} - m_i(\boldsymbol{\alpha}))^2}{2\sigma_{i,m}^2} \Bigg\} L_{m,o} \\ - \frac{24}{2} \ln(2\pi) - \frac{1}{2} \sum_{i=1}^{24} \ln(\sigma_{i,s}^2) - \sum_{i=1}^{24} \frac{(s_{i,o} - s_i(\boldsymbol{\alpha}))^2}{2\sigma_{i,s}^2} \Bigg\} L_{s,o} \quad (2)$$

17 This formulation thus explicitly takes into consideration long-term biases in
 18 the Tb average ($L_{m,o}$ [-]) and the Tb variability ($L_{s,o}$ [-]) and is derived from

1 a classical Gaussian likelihood function:

$$\begin{aligned}
p(\mathbf{m}_o, \mathbf{s}_o | \boldsymbol{\alpha}) &= \prod_{i=1}^{24} \left[\frac{1}{\sqrt{2\pi\sigma_{i,m}^2}} \exp \left(-\frac{(m_{i,o} - m_i(\boldsymbol{\alpha}))^2}{2\sigma_{i,m}^2} \right) \right] \\
&\cdot \prod_{i=1}^{24} \left[\frac{1}{\sqrt{2\pi\sigma_{i,s}^2}} \exp \left(-\frac{(s_{i,o} - s_i(\boldsymbol{\alpha}))^2}{2\sigma_{i,s}^2} \right) \right] \quad (3)
\end{aligned}$$

2 where $\sigma_{i,m}$ and $\sigma_{i,s}$ denote the (ensemble) standard deviations of the residual
3 differences between the observed and simulated values of the long-term Tb
4 averages and standard deviations, respectively. These standard deviations
5 provide a lumped description of model, input data, and observation errors.

6 3.3. Likelihood, Objective Function and Algorithm Settings

7 The design of the DREAM_(ZS) likelihood function L (Eq. 2) and the PSO
8 objective function J (Eq. B.1) warrants further discussion. As discussed
9 above, we sample the climatological, or long-term, Tb averages and standard
10 deviations over multiple incidence angles, polarizations and overpass times
11 (that is, 2×24 observations, $i = 1, \dots, 24$) per location, rather than a series
12 of observations at multiple times. The long-term Tb averages and standard
13 deviations could also be interpreted as ‘summary statistics’ or ‘signatures’
14 of the system, and hence our approach has elements in common with the
15 diagnostic model evaluation procedure presented in Vrugt and Sadegh (2013).

16 The variables $\sigma_{i,m}$ and $\sigma_{i,s}$ measure the (ensemble) standard deviation of
17 the residual differences between the observed and simulated long-term Tb
18 averages and standard deviations, respectively, for each observation i . The
19 residual errors are assumed to have a zero mean and include both SMOS
20 observation and simulation errors, due to, e.g., inaccurate soil moisture, tem-

1 perature or vegetation characteristics. These $\sigma_{i,m}$ and $\sigma_{i,s}$ statistics trade off
 2 errors in the long-term Tb averages against those of the long-term Tb stan-
 3 dard deviations (as well as deviations from the prior parameter constraints).
 4 Since only one sample is available for each observation, it is impossible to
 5 estimate individual $\sigma_{i,m}$ - and $\sigma_{i,s}$ -values. Therefore, $\sigma_{i,m}$ and $\sigma_{i,s}$ include a
 6 homoscedastic term (σ_m , σ_s) and a heteroscedastic factor w_i to account for
 7 the robustness of the diagnosed long-term Tb averages and standard devia-
 8 tions, i.e. $\sigma_{i,m}^2 = w_i \sigma_m^2$ and $\sigma_{i,s}^2 = w_i \sigma_s^2$. The homoscedastic term is identical
 9 for all 24 observations, and we either set σ_m and σ_s to a default value of 1 K
 10 (De Lannoy et al., 2013), or alternatively we estimate σ_m and σ_s jointly with
 11 the RTM parameters (see section 3.1). The weights are given by $w_i = \frac{N}{N_i}$,
 12 where N_i denotes the number of data points in time that contribute to a
 13 particular long-term Tb average (or standard deviation), and N signifies the
 14 average contributing number of time steps across all 24 observations. These
 15 weights are typically close to 1 and assign somewhat more (less) weight to
 16 climatological Tb differences that are based on more (fewer) individual data
 17 points in the different 1-year data time series that underlie the 24 different
 18 observations. For example, Tb observations at low incidence angles are by
 19 design based on fewer instantaneous data points than Tb observations at
 20 high incidence angles.

21 Per grid cell, a maximum of 12,000 log-likelihood function evaluations
 22 are performed with DREAM_(ZS) using standard settings of the algorithmic
 23 variables. For PSO, we use the same algorithmic settings as reported in
 24 De Lannoy et al. (2013), except that a swarm size of 10 particles is used with
 25 a minimum of 10 and maximum of 100 iterations. The search is terminated

1 if the reduction of the objective function is smaller than 1E-5 over the last
 2 10 iterations. A total of 12 repetitions are performed, which results in a
 3 maximum of 12,000 function evaluations.

4 3.4. Posterior Parameter Distribution

5 The MAP parameter values are defined as those with the largest value
 6 for L (Eq. 2, DREAM_(ZS)) or the smallest value for J (Eq. B.1, PSO). These
 7 MAP parameter estimates will subsequently be used in the RTM that is part
 8 of the Tb assimilation system for state updating (not discussed herein). Note
 9 that these MAP values are not necessarily identical to the posterior ensemble
 10 mean of the distribution. For the DREAM_(ZS) experiments, the last 25% of
 11 the MCMC chains (3,000 samples) are used to summarize parameter uncer-
 12 tainty by calculating the standard deviation of each individual parameter.
 13 To illustrate this in more detail for one grid cell, consider Fig. 2a, which
 14 depicts the marginal distributions of the RTM parameters. We define the
 15 uncertainty as the ensemble standard deviation $stdv[\alpha] \equiv \sqrt{(\alpha - \bar{\alpha})^2}$ cen-
 16 tralized around the ensemble mean $\bar{\alpha}$, not around the MAP parameter value
 17 α_{MAP} . The notation $\bar{\cdot}$ refers to the ensemble mean. Note that the standard
 18 deviation around the MAP estimate $stdv_{MAP}[\alpha]$ can be found as a function of
 19 the centralized standard deviation $stdv[\alpha]$, i.e. $stdv_{MAP}[\alpha]^2 = stdv[\alpha]^2 + \Delta_\alpha^2$,
 20 where $\Delta_\alpha = \bar{\alpha} - \alpha_{MAP}$ is the difference between the ensemble mean and
 21 MAP parameter estimate. We found that, across the different experiments,
 22 Δ_α is usually small (see section 4.1), so that $stdv_{MAP}[\alpha] \sim stdv[\alpha]$.

3.5. Tb Simulation Performance and Uncertainty

A number of measures are used to evaluate the long-term Tb simulations and their associated uncertainty. Fig. 2b illustrates some of the terms used in this evaluation. We assess the quality of the deterministic Tb simulations with the MAP parameter estimates, using the mean-square difference (MSD [K²]) between the observed and simulated long-term Tb averages (Eq. 4) and standard deviations (Eq. 5) across the 24 different observations:

$$MSD_m = \frac{1}{24} \sum_{i=1}^{24} (m_i(\boldsymbol{\alpha}_{MAP}) - m_{i,o})^2 \quad (4)$$

$$MSD_s = \frac{1}{24} \sum_{i=1}^{24} (s_i(\boldsymbol{\alpha}_{MAP}) - s_{i,o})^2 \quad (5)$$

If the modeling errors were solely due to inaccurate parameter values and the observational terms were error-free, these metrics should be very close to zero. In practice, however, the model structure is not perfect and model inputs such as soil moisture and temperature as well as the observational terms are subject to errors. Therefore, the metrics will substantially deviate from zero and reflect the total residual simulation and observation errors. The 24 differences contributing to MSD_m are illustrated as Δ_{m_i} in Fig. 2b.

If the uncertainties are well estimated and biases between observations and simulations are constrained during the calibration, then the “actual” (MSD_m, MSD_s) and “expected” ensemble ($\sigma_{i,m}^2, \sigma_{i,s}^2$) residual Tb error variances should be equal, or their ratio should be close to 1. Note that a similar check of consistency is used to verify the prescribed observation and simulation uncertainties in data assimilation systems (Reichle et al., 2002) and for ensemble forecast verification (De Lannoy et al., 2006). The only difference is

1 that here, the mean values (i.e. the ‘M’, or mean, in MSD) are derived from
 2 multiple observations types ($i = 1, \dots, 24$), whereas in the earlier studies the
 3 mean was calculated in the time domain.

4 The above total residual Tb error lumps all long-term errors in the model,
 5 input data, and observations. The Tb simulation error due to parameter un-
 6 certainty can be isolated. To this end, we analyze an ensemble of Tb simula-
 7 tions, obtained by propagating 20 samples from the MCMC derived posterior
 8 parameter distributions through the RTM. The performance of the ensemble
 9 mean of simulated long-term Tb averages $\overline{m_i(\boldsymbol{\alpha})}$ and standard deviations
 10 $\overline{s_i(\boldsymbol{\alpha})}$ is given by:

$$MSD_{\overline{m}} = \frac{1}{24} \sum_{i=1}^{24} \left(\overline{m_i(\boldsymbol{\alpha})} - m_{i,o} \right)^2 \quad (6)$$

$$MSD_{\overline{s}} = \frac{1}{24} \sum_{i=1}^{24} \left(\overline{s_i(\boldsymbol{\alpha})} - s_{i,o} \right)^2 \quad (7)$$

11 where $\overline{\cdot}$ denotes the ensemble mean. Fig. 2b illustrates the 24 differences
 12 contributing to $MSD_{\overline{m}}$ as $\Delta_{\overline{m}_i}$. Because of the non-linear nature of the RTM
 13 and because $\overline{\boldsymbol{\alpha}}$ and $\boldsymbol{\alpha}_{MAP}$ generally differ, $MSD_{\overline{m}}$ and $MSD_{\overline{s}}$ will deviate
 14 from MSD_m and MSD_s , but the differences should be limited if the posterior
 15 RTM parameter distributions are narrow.

16 The Tb uncertainty due to parameter error is quantified by the ensemble
 17 standard deviation in simulated long-term Tb averages ($\sigma_{i,m,par}$, illustrated in
 18 Fig. 2b) and long-term Tb standard deviations ($\sigma_{i,s,par}$). The corresponding
 19 ensemble variance for each observation type i is:

$$\sigma_{i,m,par}^2 = \overline{\left(m_i(\boldsymbol{\alpha}) - \overline{m_i(\boldsymbol{\alpha})} \right)^2} \quad (8)$$

$$\sigma_{i,s,par}^2 = \overline{\left(s_i(\boldsymbol{\alpha}) - \overline{s_i(\boldsymbol{\alpha})} \right)^2} \quad (9)$$

1 Averaged across 24 observations, this results in a mean ensemble spread of:

$$ME nSp_{m,par} = \frac{1}{24} \sum_{i=1}^{24} \sigma_{i,m,par}^2 \quad (10)$$

$$ME nSp_{s,par} = \frac{1}{24} \sum_{i=1}^{24} \sigma_{i,s,par}^2 \quad (11)$$

2 The above metrics are primarily important to quantify the skill of the
 3 posterior Tb simulations. For comparison, we also calculate the MSD_m and
 4 MSD_s using prior RTM parameters (Table 1). For clarity, all metrics in this
 5 section are expressed as variances [K^2], but in the results below we present
 6 Tb uncertainties as the corresponding standard deviations [K].

7 4. Results

8 4.1. RTM Parameters and Their Uncertainty

9 In this section, we analyze the MAP values of the microwave surface
 10 roughness $\langle h \rangle$, the vegetation optical depth $\langle \tau \rangle$, the scattering albedo
 11 ω , and their posterior uncertainty ($stdv[.]$). As will be shown below, the
 12 DREAM_(ZS) case D_σ should be considered as benchmark throughout the
 13 paper, because of the statistical rigor of the sampled posterior. Fig. 3 shows
 14 maps of the prior parameter values and the MAP estimates derived from
 15 experiment cases P, D and D_σ (Table 1). The spatially averaged posterior
 16 parameter values are very similar for all 3 cases, with a microwave roughness
 17 $\langle h \rangle$ around 0.72 ± 0.5 [-], a vegetation optical depth $\langle \tau \rangle$ of 0.26 ± 0.15 [-]
 18 and a scattering albedo ω of 0.09 ± 0.07 [-], where the values after the \pm sign
 19 measure the spatial standard deviation and reflect the variability of the MAP
 20 parameters across the spatial domain. Note that these latter values should

1 not be confused with uncertainty estimates. Compared to the prior values,
2 $\langle h \rangle$ is generally larger for grassland, $\langle \tau \rangle$ is smaller for forests and ω is
3 larger for all vegetation classes except grassland (details per vegetation class
4 not shown; these findings are similar to those of De Lannoy et al. (2013)). The
5 spatial patterns for the 3 experiments are also very similar, with exceptions
6 discussed below. Moreover, Fig. 3 suggests that MAP values derived with
7 the PSO algorithm closely match those of DREAM_(ZS).

8 Fig. 4 shows the ensemble parameter uncertainty for cases D and D _{σ} .
9 Maps with RTM parameter uncertainty estimates for PSO (obtained as in
10 De Lannoy et al. (2013)) are not shown, because they are statistically invalid
11 and significantly larger than those derived with DREAM_(ZS). The prior pa-
12 rameter uncertainty is also not shown, because it is spatially uniform. Table 1
13 indicates that the relative uncertainty in the literature-based prior param-
14 eters exceeds 100% of α_0 . In contrast, the relative uncertainties for case D
15 (Fig. 4a-c) are less than 10% of the MAP parameter value and substantially
16 smaller than the spatial variability in the MAP values. For case D _{σ} , the
17 relative uncertainties are ranging up to 25% of the MAP values: for $\langle h \rangle$,
18 the spatially averaged uncertainty is 0.10 [-], for $\langle \tau \rangle$ 0.04 [-] and for ω
19 0.02 [-]. The uncertainty in $\langle h \rangle$ typically increases with more complex
20 terrain and is smallest in the cropped region southwest of the Great Lakes.
21 The uncertainty of $\langle \tau \rangle$ is largest in the forested Appalachian mountains
22 where the highest MAP values of $\langle \tau \rangle$ are found. On the contrary, ω is
23 best defined in this area and uncertainties in ω increase in the dry Western
24 mountain ranges. The $\langle h \rangle$ -values exhibit more uncertainty where either
25 the uncertainty in ω (Fig. 4e) or $\langle \tau \rangle$ (Fig. 4f) is larger. The global mean

absolute differences between the MAP and ensemble mean parameter values (not shown; discussed in section 3.4) for $\langle h \rangle$, $\langle \tau \rangle$ and ω are $\Delta\alpha=0.07$, 0.02 and 0.01 [-], respectively, for case D_σ , and $\Delta\alpha=0.02$, 0.01 and 0.00 [-], respectively, for case D.

In summary, both DREAM_(ZS) cases D and D_σ provide MAP parameter values that are very similar and in close agreement with the PSO estimates (Fig. 3). The DREAM_(ZS) derived posterior parameters appear well defined with relative uncertainties that are less than 25% of the MAP values. It will be shown below that the uncertainty estimates of case D_σ – unlike those of case D – are consistent with the sample root-mean-square difference (RMSD) between long-term Tb observations and simulations.

4.2. Residual Tb Error Standard Deviations

To analyze the differences between cases D and D_σ , or specifically the effect of estimating the residual error standard deviations of the long-term Tb averages (σ_m) and standard deviations (σ_s), we list the domain-averaged MAP parameter values and their associated uncertainties in Table 2 for all experiment cases. In addition, Fig. 5 depicts the results for different vegetation classes. As discussed above, cases D and D_σ return similar MAP values for the RTM parameters with some local exceptions, such as for example for ω over cropland (Fig. 5e). The estimated posterior RTM parameter uncertainty increases about 2 - 3 times, when σ_m and σ_s are included in the parameter estimation (i.e., case D_σ). For case D_σ , the domain-averaged values are $\sigma_m = 3.5$ K and $\sigma_s = 2.3$ K (Table 2), whereas case D uses default values of $\sigma_m = \sigma_s = 1$ K. The σ_m and σ_s estimates are not defined for the simulations with prior parameters.

1 The residual error standard deviations σ_m and σ_s are estimated to be
 2 larger than the initially imposed 1 K because of significant Tb observation
 3 and simulation errors that are not due to parameter errors. It is necessary
 4 to use this increased residual Tb error in the RTM parameter estimation to
 5 ensure parameters that do not, or at least only minimally, compensate for
 6 errors other than parameter error (i.e. error in geophysical input fields, aux-
 7 iliary information, RTM structure or in observations). The local differences
 8 between the MAP values for cases D and D_σ are thus explained by how much
 9 the parameter values are forced to compensate for errors other than param-
 10 eter error. As will be shown below, this can result in suboptimal parameter
 11 values for case D. The uncertainty in the RTM parameters for D_σ is larger,
 12 because we allow for more realistic Tb simulation and observation error, re-
 13 sulting in a wider range of acceptable RTM parameters. From a Bayesian
 14 perspective, the observation likelihood function becomes wider (larger σ_m
 15 and σ_s) and thus the posterior parameter probability function is expected to
 16 be wider.

17 The value of σ_m and its posterior uncertainty are largest in cropped re-
 18 gions (Fig. 5g) where residual Tb errors are dominated by less skillful model
 19 simulations. This is to be expected because irrigation is not simulated and
 20 the climatological LAI estimates do not account for interannual crop ro-
 21 tations. The parameters should not compensate for these errors, and the
 22 default values of $\sigma_m = \sigma_s = 1$ K make cases D and P vulnerable to subopti-
 23 mal solutions. For example, the relatively larger differences between D and
 24 D_σ for σ_m over cropland result in larger differences in their respective MAP
 25 values of ω (Fig. 5i). For forests, $\sigma_s = 1$ K appears to be a good estimate

1 (Fig. 5i) because the variability in Tb is expected to be low due to vegetation
2 attenuation.

3 Both the MAP values and the uncertainties for σ_m are always larger
4 than those derived for σ_s . One of the reasons for the higher σ_m values are
5 the opposite signs in the biases for the long-term Tb averages at morning
6 and evening overpasses, which cannot be mitigated with time-invariant RTM
7 parameters. These biases are due to sensor error and modeled temperature
8 errors as discussed in De Lannoy et al. (2013). In a separate exercise (not
9 shown herein), we verified that the σ -values absorb biases in geophysical
10 fields: by re-scaling the soil moisture the RMSD and σ -values are jointly
11 reduced.

12 Based on these new insights, we currently use globally uniform values of
13 $\sigma_m = 3.5$ K and $\sigma_s = 2.3$ K for the SMOS-based calibration of the RTM pa-
14 rameters that will be needed to generate the forthcoming SMAP L4_SM prod-
15 uct. This represents an advance over the mere guesses of $\sigma_m = \sigma_s = 1.0$ K
16 used in earlier work (De Lannoy et al., 2013). The resulting parameter es-
17 timates optimally mimic the true relationship between land surface fields
18 and Tb. However, larger σ -values likely increase the error of the Tb simu-
19 lations, because the model parameters will compensate less for local biases
20 in, for example, soil moisture. These issues will be addressed through model
21 development and bias correction inside the data assimilation system.

22 4.3. Posterior Tb Simulation Performance

23 The quality of the estimated MAP parameters can be measured by the
24 skill of the corresponding Tb simulations. Fig. 6 shows the misfit between
25 observed and MAP simulated long-term Tb averages and standard deviations

1 ($RMSD_m$, $RMSD_s$, square root of Eq. 4 and 5) across the 24 observations
 2 for the calibration and evaluation period, averaged per vegetation class. Ta-
 3 ble 2 lists the domain-averaged $RMSD_m$ and $RMSD_s$ values during the
 4 calibration period. The prior parameters result in $RMSD_m=12.9$ K and
 5 $RMSD_s=3.2$ K. The skill is greatly improved after parameter estimation
 6 and very similar for cases P, D and D_σ ($RMSD_m$ around 3 K and $RMSD_s$
 7 around 2.5 K), which is not surprising given that the three different cases gen-
 8 erate similar MAP parameter estimates. During the evaluation period, the
 9 $RMSD_m$ increases up to 8 K for cropland (Fig. 6b), and the $RMSD_s$ reaches
 10 values of 5 K for cropland in the evaluation year (Fig. 6d). Cropland exhibits
 11 the largest errors, because of known simulation errors (see above). The larger
 12 errors in the evaluation period suggest that Tb simulations for interannually
 13 varying agricultural areas will be compromised when using time-invariant
 14 RTM parameters and climatological vegetation information.

15 Table 2 also shows the $RMSD_{\overline{m}}$ and $RMSD_{\overline{s}}$ (square root of Eqs. 6 and
 16 7) for the ensemble mean of the Tb simulations $\overline{m_i(\boldsymbol{\alpha})}$ and $\overline{s_i(\boldsymbol{\alpha})}$. Ensembles
 17 are generated by sampling the posterior parameter distribution. The lim-
 18 ited uncertainty in the posterior parameters results in comparable ensemble
 19 mean and deterministic Tb simulations, so that $RMSD_m \sim RMSD_{\overline{m}}$ and
 20 $RMSD_s \sim RMSD_{\overline{s}}$.

21 4.4. Actual and Expected Tb Simulation Errors

22 If the parameter estimation procedure is statistically consistent, then the
 23 actual Tb errors ($RMSD_m$, $RMSD_s$, section 4.3) and the estimated residual
 24 Tb errors (σ_m , σ_s , section 4.2) should be of similar magnitude. Indeed, for
 25 case D_σ the $RMSD_m$ values during the calibration period (Fig. 6a) and

1 the σ_m estimates (Fig. 5g) show similar variations across vegetation classes.
 2 Likewise, the $RMSD_s$ values (Fig. 6b) and the σ_s estimates (Fig. 5i) show
 3 similar variations across vegetation classes. For case D, however, the same
 4 is not true. Moreover, Table 2 suggests that the domain-averaged ratio of
 5 $RMSD_m$ and $RMEnSp_m$ for the long-term average Tb is 2.9 for case D
 6 and 1.0 for case D_σ . Similarly, the domain-averaged ratio of $RMSD_s$ and
 7 $RMEnSp_s$ is 2.5 for case D and 1.0 for case D_σ . Adequate results are thus
 8 only found for case D_σ by estimating σ_m and σ_s , whereas case D falls short
 9 with respect to these metrics. Note that during the evaluation period (not
 10 shown), the ratios always exceed 1, because of an increased $RMSD_m$ and
 11 $RMSD_s$. Nevertheless, case D_σ is still more consistent than case D.

12 Table 2 also lists the square root of the mean simulation uncertainty due
 13 to posterior parameter error, i.e. $RMEnSp_{m,par}$ and $RMEnSp_{s,par}$ (square
 14 root of Eqs. 10 and 11). These metrics are not shown for the prior and PSO
 15 cases, because those parameter distributions are not sampled adequately.
 16 After parameter estimation, $RMEnSp_{m,par}=0.3$ K and $RMEnSp_{s,par}=0.2$ K
 17 for case D, whereas $RMEnSp_{m,par}=1.0$ K and $RMEnSp_{s,par}=0.4$ K for case
 18 D_σ . The uncertainty associated with the posterior parameter values is only
 19 a small fraction of the total residual error, that is, $RMEnSp_{par} \ll \sigma$ after
 20 parameter estimation. This gives us confidence that the RTM parameters
 21 have been estimated reliably and that other errors dominate the residual
 22 errors or biases in the long-term Tb averages and standard deviations.

23 Fig. 7 illustrates that the estimated ensemble residual Tb error standard
 24 deviations of case D_σ are consistent with the actual residuals between Tb
 25 observations and simulations. Specifically, Fig. 7a shows the SMOS observed

$m_{i,o}$ as dots and the GEOS-5 simulated $m_i(\boldsymbol{\alpha}_{MAP})$ for H-polarized Tb at 6 angles at morning overpasses for cases D and D_σ , as horizontal dashes. The error bars around the Tb simulations reflect twice the lumped uncertainty due to ensemble simulation and observation errors, or $2\sigma_{i,m}$ and $2\sigma_{i,s}$. The plotted results are an average over the entire study domain for the morning overpasses. Fig. 7b shows the same for V-polarized Tb, and Figs. 7c and d provide this information for the long-term Tb standard deviations. The results for the 24 observations derived from evening overpasses are very similar and not shown.

If the uncertainty treatment underlying the parameter estimation is consistent, then 95% of the observations are expected to fall within $2\sigma_m$ or $2\sigma_s$ around both sides of the Tb simulations (assuming Gaussian distributions). The error bars for case D_σ envelop all observations, whereas the error bars for case D enclose less than half of all observations. Fig. 7 also explains the nature of the residual misfit. With the exception of the 57.5° -angle, the Tb simulations $m_i(\boldsymbol{\alpha}_{MAP})$ for morning overpasses consistently underestimate the SMOS-observed $m_{i,o}$ for H-polarization (Fig. 7a) and randomly deviate at V-polarization (Fig. 7b). In contrast, the evening simulations $m_i(\boldsymbol{\alpha}_{MAP})$ slightly overestimate the SMOS-observed $m_{i,o}$ at H-polarization (not shown herein, see De Lannoy et al. (2013)). The SMOS-observed $s_{i,o}$ is always larger than the simulated $s_i(\boldsymbol{\alpha}_{MAP})$. This is probably dominated by observation noise, but could also be attributed to an underestimation of the variability in the Tb simulations. For example, an increase in the RTM parameter h not only compensates for a cold Tb bias but simultaneously also reduces the Tb variability. Fig. 7 thus clearly illustrates why the uncertainty estimates

1 obtained from case D_σ are superior.

2 4.5. Convergence and Computational Cost

3 The effectiveness of the parameter sampling algorithm is also measured
4 by the convergence of the sampled posterior distribution to the target distri-
5 bution. For DREAM_(ZS), the potential scale reduction factor \sqrt{R} by Gelman
6 and Rubin (1992) should be near 1 to inspire confidence that the differ-
7 ent MCMC chains have converged to the appropriate target distribution.
8 The target variance is estimated based on the variances within and between
9 chains up to the current iteration. The \sqrt{R} metric measures by which scale
10 the posterior distribution at the current iteration would shrink as the num-
11 ber of MCMC iterations would go to infinity to reach the target distribution.
12 Fig. 8 shows the evolution of \sqrt{R} , averaged over all estimated parameters
13 and across the study domain. Initially, the values of \sqrt{R} are rather large
14 (due to random initial sampling) before they settle down and reach values
15 close to 1.

16 Finally, we report that the derivation of the posterior distributions re-
17 quires approximately 225 seconds of wall time for a single grid cell using
18 DREAM_(ZS) on a state-of-the-art, single-processor computing platform. For
19 global applications that involve $10^5 - 10^6$ grid cells, exploration of the poste-
20 rior distribution may be too costly. Yet, if our main interest is in the MAP
21 values, and we thus only aim at obtaining a long-term unbiased modeling
22 system, then PSO or DREAM_(ZS) are both viable options.

1 5. Conclusions

2 Accurate estimates of microwave RTM parameters for large-scale L-band
3 applications are difficult to obtain. The available parameter estimates are
4 generally based on small-scale field experiments and often come without any
5 estimate of their uncertainty. This complicates radiative transfer modeling
6 for both the forward simulation of L-band Tb over land and the retrieval of
7 soil moisture based on Tb observations. Our study presents reliable, large-
8 scale estimates of RTM parameter values and their uncertainty, based on the
9 combination of long-term (climatological) information from satellite obser-
10 vations and land model simulations. These estimated parameter values can
11 serve to limit biases in large-scale Tb simulations and soil moisture retrievals.
12 The estimates of parameter uncertainties help to derive better informed un-
13 certainties for long-term (climatological) Tb simulations and retrievals, and
14 they aid in the development of land surface data assimilation systems. This
15 is a significant advancement compared to the common practice of relying on
16 literature-based parameter values or parameter calibration approaches that
17 do not use Bayesian methods with proper characterization of model and ob-
18 servation errors.

19 More specifically, the present paper expands earlier research reported in
20 De Lannoy et al. (2013) to derive time-invariant RTM parameters and their
21 uncertainties using observations of the long-term average Tb and the long-
22 term Tb standard deviation obtained from SMOS data. The overall objec-
23 tive is to optimize GEOS-5 Tb simulations prior to sequential assimilation
24 of SMOS or SMAP Tb data, such as planned for the SMAP L4_SM product
25 (Reichle et al., 2012), and to examine the uncertainties involved in the opti-

1 mization. Per grid cell, 48 observations of the long-term SMOS Tb averages
 2 and standard deviations are constructed, i.e. Tb averages and standard de-
 3 viations for 24 different combinations of 6 incidence angles, 2 polarizations
 4 and 2 overpass times. The differences between the observed and GEOS-5
 5 simulated long-term Tb averages and standard deviations are minimized (as
 6 opposed to minimizing differences between Tb observations and simulations
 7 in the time domain) and used along with the prior parameter information to
 8 derive posterior parameter estimates.

9 The full posterior distribution of RTM parameters is derived using MCMC
 10 simulation with the DREAM_(ZS) algorithm. To our knowledge, this is the
 11 first large-scale application of the DREAM_(ZS) algorithm for the estimation
 12 of RTM parameters and their underlying uncertainty. The results serve as
 13 a benchmark to verify the results from simpler parameter optimization al-
 14 gorithms, such as for example PSO. Such algorithms may be desirable for
 15 global-scale operational applications that rely on evolving modeling systems
 16 in need of frequent re-calibrations, without a need to fully sample the poste-
 17 rior parameter distribution.

18 Firstly, we verified that the MAP RTM-parameter values derived from
 19 converged posterior distributions with DREAM_(ZS) can be approximated by
 20 a simpler optimization algorithm (PSO), which corroborates our earlier re-
 21 search (De Lannoy et al., 2013). The MAP parameters will later be used
 22 in an RTM that is part of a data assimilation system to ensure minimally
 23 biased Tb simulations. Secondly, we obtained reliable estimates of parameter
 24 uncertainty with DREAM_(ZS), which cannot be obtained with PSO. The rel-
 25 ative parameter uncertainties are generally less than 25% of the MAP value

1 for $\langle h \rangle$, $\langle \tau \rangle$ and ω , when including an estimation of the residual (obser-
2 vation and simulation) error standard deviations (σ_m , σ_s) of the long-term
3 Tb averages and standard deviations. This relative parameter uncertainty is
4 substantially reduced compared to the relative uncertainty in literature-based
5 parameter values, which easily exceeds 100%.

6 The third objective of this paper was to quantify the importance of pa-
7 rameter and other errors on long-term Tb simulations. The MAP parameters
8 reduce the RMSD in long-term Tb averages and standard deviations to 3.4 K
9 and 2.3 K, respectively. This is a reduction by 74% and 30% compared to
10 results with prior parameters. Of this total *posterior* Tb uncertainty, only
11 1 K and 0.4 K is associated with the posterior uncertainty in the parameter
12 values.

13 The actual RMSD in the long-term Tb averages and standard deviations
14 is matched by the estimated ensemble residual Tb error standard deviation
15 ($\sigma_m=3.5$ K, $\sigma_s=2.3$ K, assumed homoscedastic), obtained with DREAM_(ZS)
16 case D _{σ} . In other words, the joint estimation of RTM parameters and σ_m and
17 σ_s results in a balance between actual and expected errors in Tb simulations,
18 and in statistically adequate parameter values and uncertainty estimates.
19 The prior estimate of 1 K for σ_m and σ_s , used in case P, case D and De Lan-
20 noy et al. (2013) was thus too low, except for σ_s over forests which exhibit
21 limited Tb variability due to vegetation attenuation. The largest σ_m -values
22 are found in cropped regions where the lumped residual Tb errors are prob-
23 ably dominated by errors in geophysical fields (e.g. vegetation, soil moisture
24 and temperature) that constitute important inputs to the Tb simulations.

25 In practice, these findings prompt us to revise the global RTM calibra-

tion of De Lannoy et al. (2013) with improved residual error estimates in preparation for the SMAP L4_SM product. The revised approach ensures parameters that optimally describe the true relationship between land surface fields and Tb, while minimally compensating for potential biases in any of these fields. Furthermore, the findings of the present study indicate that, *after* RTM parameter estimation, the residual climatological uncertainties reside in Tb observation error and GEOS-5 Tb simulation error that is not, or only to a limited extent, related to RTM parameter uncertainty.

In summary, the Bayesian inference of the posterior distribution of the RTM parameters ensures reliable Tb simulations with GEOS-5. The MAP parameter estimates guarantee a Tb assimilation system with limited biases and a realistic connection between Tb and land surface fields such as soil moisture and temperature. Furthermore, the DREAM_(ZS) algorithm reveals the importance of observation error and simulation error that cannot be explained by the RTM parameters. The posterior Tb uncertainties in this study pertain to long-term Tb averages and standard deviations and are thus indicative of biases (i.e. long-term errors). These residual biases will be addressed by model refinement and bias mitigation during the assimilation of satellite-observed Tb data.

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1 **Appendix A. Radiative Transfer Model (RTM)**

2 A diagnostic zero-order (tau-omega) microwave RTM is used to simulate
 3 L-band Tb at the top of the atmosphere ($Tb_{TOA,p}$ [K]). The $Tb_{TOA,p}$ at
 4 polarization $p = [H, V]$ (horizontal or vertical) is a combination of (i) soil
 5 emission, possibly attenuated by vegetation, (ii) vegetation emission, possibly
 6 reflected by the soil, and (iii) atmospheric effects:

$$\begin{aligned} Tb_{TOV,p} &= T_s(1 - r_p)A_p + T_c(1 - \omega_p)(1 - A_p)(1 + r_pA_p) \\ &\quad + Tb_{ad,p}r_pA_p^2 \end{aligned} \quad (A.1)$$

$$Tb_{TOA,p} = Tb_{au,p} + \exp(-\tau_{atm,p})Tb_{TOV,p} \quad (A.2)$$

7 where $Tb_{TOV,p}$ [K] is the top of vegetation Tb, T_s [K] is the surface soil tem-
 8 perature, T_c [K] is the canopy temperature (assumed equal to T_s), $Tb_{ad,p}$ [K]
 9 and $Tb_{au,p}$ [K] are the downward and upward atmospheric radiation, A_p [-] is
 10 the vegetation attenuation, $\exp(-\tau_{atm,p})$ [-] is the atmospheric attenuation,
 11 $\tau_{atm,p}$ [-] is the atmospheric optical depth, r_p [-] is the rough surface reflec-
 12 tivity, and ω_p [-] is the scattering albedo. The atmospheric contributions
 13 ($Tb_{ad,p}$, $Tb_{au,p}$ and $\exp(-\tau_{atm,p})$) are described by Pellarin et al. (2003). The
 14 rough surface reflectivity r_p [-] is derived from the smooth surface reflectivity
 15 R_p [-] following Choudhury et al. (1979) and Wang and Choudhury (1981):

$$r_p = (Q R_q + (1 - Q)R_p) \exp(-h) \cos^{Nr_p}(\theta) \quad (A.3)$$

16 where Q [-] is the polarization mixing ratio and typically set to 0 for L-
 17 band (Kerr and Njoku, 1990), θ [rad] is the incidence angle, h [-] is the
 18 roughness parameter accounting for dielectric properties that vary at the sub-
 19 wavelength scale, Nr_p [-] is the angular dependence, and $q = V$ for $p = H$

1 and vice versa. The smooth surface reflectivity R_p [-] is given by the Fresnel
2 equations as a function of the dielectric constant, which itself depends on soil
3 moisture, temperature, texture, incidence angle and wavelength. We select
4 the Wang and Schmugge (1980) soil dielectric mixing model for this study.
5 The results with this model are similar to what is obtained with the Mironov
6 et al. (2004) model, and both are in a better agreement with the SMOS data
7 than the Dobson et al. (1985) model. We include the dependence of h on
8 soil moisture (SM [$\text{m}^3.\text{m}^{-3}$]) through a stepwise linear expression (adapted
9 from the proposed SMOS soil moisture retrieval algorithm (CESBIO et al.,
10 2011; Kerr et al., 2012)):

$$h = \begin{cases} h_{max} & \text{if } SM \leq wt \\ h_{max} + \frac{h_{min}-h_{max}}{poros-wt} (SM - wt) & \text{if } wt < SM \leq poros \end{cases} \quad (\text{A.4})$$

11 where $poros$ [$\text{m}^3.\text{m}^{-3}$] and wt [$\text{m}^3.\text{m}^{-3}$] are the porosity and transition soil
12 moisture, respectively. The latter is modeled as $wt = 0.48.wp + 0.165$ (Wang
13 and Schmugge (1980)) where wp [$\text{m}^3.\text{m}^{-3}$] is the wilting point.

14 The vegetation attenuation A_p [-] is based on the Jackson and Schmugge
15 (1991) vegetation opacity model:

$$A_p = \exp\left(-\frac{\tau_p}{\cos \theta}\right), \text{ with} \quad (\text{A.5})$$

$$\tau_p = b_p VWC = b_p LEWT LAI \quad (\text{A.6})$$

16 where τ_p [-] is the nadir vegetation optical depth, which is a function of
17 a vegetation structure parameter b_p [-] and the vegetation water content
18 (VWC) [$\text{kg}.\text{m}^{-2}$]. The latter is modeled as the product of LAI [$\text{m}^2.\text{m}^{-2}$]
19 and the leaf equivalent water thickness ($LEWT$) [$\text{kg}.\text{m}^{-2}$].

1 Appendix B. Particle Swarm Optimization (PSO)

2 The PSO algorithm (Kennedy and Eberhart, 1995) is a global search
 3 method that uses a dynamic swarm of particles to explore the parameter
 4 space. The best position of each individual particle (cognitive aspect) and
 5 of the entire swarm (social aspect) are used to guide the particles towards
 6 the optimal solution. The iterative swarm search is performed in several
 7 independent repetitions to account for sampling variability.

8 The fitness of each parameter combination in the swarm is measured by
 9 an integrated ‘cost’ or ‘objective function’ J [-] that measures the distances
 10 between the observed and simulated long-term Tb averages ($J_{m,o}$ [-]) and
 11 standard deviations ($J_{s,o}$ [-]). To make sure that the estimated parameter
 12 values honor the prior information (as used in DREAM_(ZS)), we also include
 13 a penalty term that quantifies deviations of the parameters from their a
 14 priori expected values (J_α [-]). This results in the following definition of the
 15 objective function to be minimized:

$$\begin{aligned}
 J = & \left. \sum_{i=1}^{24} \frac{(m_{i,o} - m_i(\boldsymbol{\alpha}))^2}{2\sigma_{i,m}^2} \right\} J_{m,o} \\
 & + \left. \sum_{i=1}^{24} \frac{(s_{i,o} - s_i(\boldsymbol{\alpha}))^2}{2\sigma_{i,s}^2} \right\} J_{s,o} \\
 & + \left. \sum_{k=1}^{N_\alpha} \frac{(\alpha_{0,k} - \alpha_k)^2}{2\sigma_{\alpha_{0,k}}^2} \right\} J_\alpha
 \end{aligned} \tag{B.1}$$

16 where N_α signifies the number of simultaneously estimated parameters. Note
 17 that Eq. B.1 is a polished version of Eq. 6 in De Lannoy et al. (2013), which
 18 differs in the presentation of the weight factors and constants. This formu-
 19 lation is essentially similar to the definition of the posterior density used in

1 DREAM_(ZS). The main difference is that PSO handles the prior information
2 of the parameters explicitly as penalty term J_α in the objective function,
3 whereas in DREAM_(ZS), the prior parameter distribution is handled inde-
4 pendently from the likelihood function by application of Bayes law. Both
5 methods should thus find the same “best” parameter values.

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Table 1: Parameters selected (X) for calibration in different experiment cases (P, D, D_σ) with an indication of the allowed parameter range ($[\alpha_{min}, \alpha_{max}]$) and the prior estimate for each IGBP vegetation class within the study domain. $\Delta h \equiv h_{max} - h_{min}$, $\Delta b \equiv b_V - b_H$.

		h_{min} [-]	Δh [-]	ω [-]	b_H [-]	Δb [-]	σ_m [K]	σ_s [K]
Case P (PSO)		X	X	X	X	X	-	-
Case D ($DREAM_{(ZS)}$)		X	X	X	X	X	-	-
Case D_σ ($DREAM_{(ZS)}$)		X	X	X	X	X	X	X
α_{min}		0	0	0	0	-0.15	1E-5	1E-5
α_{max}		2.0	1.0	0.3	0.7	0.15	60	40
ENF	Evergreen Needleleaf Forest	1.2	0	0.05	0.33	0	1	1
DBF	Deciduous Broadleaf Forest	1	0	0.05	0.33	0	1	1
MXF	Mixed Forest	1.3	0	0.05	0.33	0	1	1
CSH	Closed Shrublands	0.7	0	0.05	0.3	0	1	1
OSH	Open Shrublands	0.7	0	0.05	0.3	0	1	1
WSV	Woody Savannas	0.7	0	0.05	0.3	0	1	1
GRS	Grasslands	0.1	0	0.05	0.2	0	1	1
CRP	Croplands	0.5	0	0.05	0.15	0	1	1
CRN	Crop and Natural Vegetation	0.7	0	0.05	0.15	0	1	1

Table 2: Domain-averaged parameter values and their uncertainty $stdv[.]$ for the prior distributions and the posterior distributions obtained with cases P, D and D_σ . The RMSD values for MAP and ensemble mean Tb simulations (square root of Eqs. 4-7) and the ensemble standard deviations in Tb simulations due to parameter errors (square root of Eqs. 10-11) are averaged across 24 long-term Tb observations and calculated for the calibration period.

	Prior	P	D	D_σ
$\langle h \rangle [-]$	0.59	0.72	0.73	0.72
$\langle \tau \rangle [-]$	0.35	0.26	0.26	0.25
$\omega [-]$	0.05	0.09	0.09	0.08
σ_m [K]	-	1.00	1.00	3.45
σ_s [K]	-	1.00	1.00	2.26
$stdv[\langle h \rangle] [-]$	0.63	-	0.04	0.10
$stdv[\langle \tau \rangle] [-]$	0.27	-	0.02	0.04
$stdv[\omega] [-]$	0.09	-	0.01	0.02
$stdv[\sigma_m]$ [K]	-	-	-	0.78
$stdv[\sigma_s]$ [K]	-	-	-	0.51
$RMSD_m$ [K]	12.90	2.98	2.90	3.41
$RMSD_s$ [K]	3.21	2.57	2.45	2.25
$RMSD_{\overline{m}}$ [K]	-	-	2.94	3.37
$RMSD_{\overline{s}}$ [K]	-	-	2.51	2.42
$RME nSp_{m,par}$ [K]	-	-	0.29	1.00
$RME nSp_{s,par}$ [K]	-	-	0.15	0.39

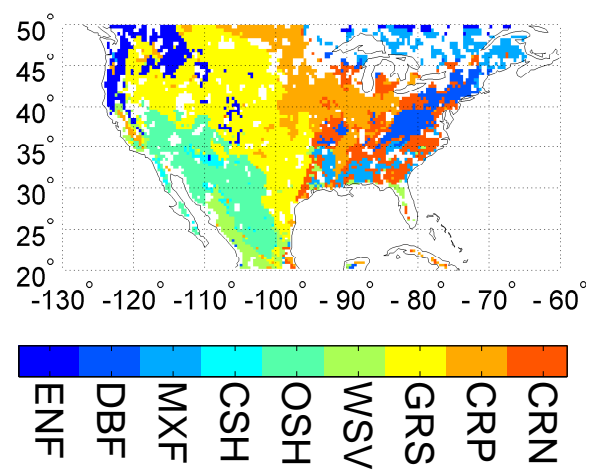


Figure 1: Study domain with indication of the dominant IGBP vegetation classes.

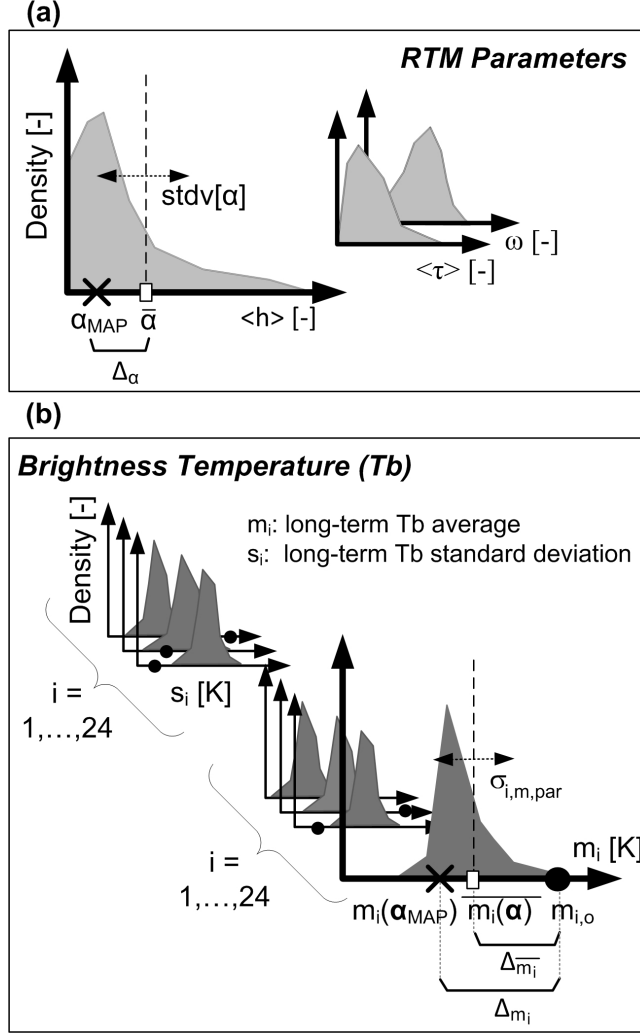


Figure 2: Illustration of marginal distributions for (a) RTM parameters and (b) Tb simulations at a single grid cell. Crosses (\times) indicate the MAP estimates, the vertical dashed lines and white squares indicate the ensemble mean posterior estimates, and horizontal dotted arrows indicate one standard deviation uncertainty around the ensemble mean. The performance of the Tb simulations is quantified by comparing either the MAP ($m_i(\alpha_{MAP})$, $s_i(\alpha_{MAP})$) or the ensemble mean ($\bar{m}_i(\alpha)$, $\bar{s}_i(\alpha)$) simulations against (black dots) 24 observed values ($m_{i,o}$, $s_{i,o}$) with $i = 1, \dots, 24$. The differences Δ_{m_i} and $\Delta_{\bar{m}_i}$ contribute to MSD_m (Eq. 4) and $MSD_{\bar{m}}$ (Eq. 6), respectively.

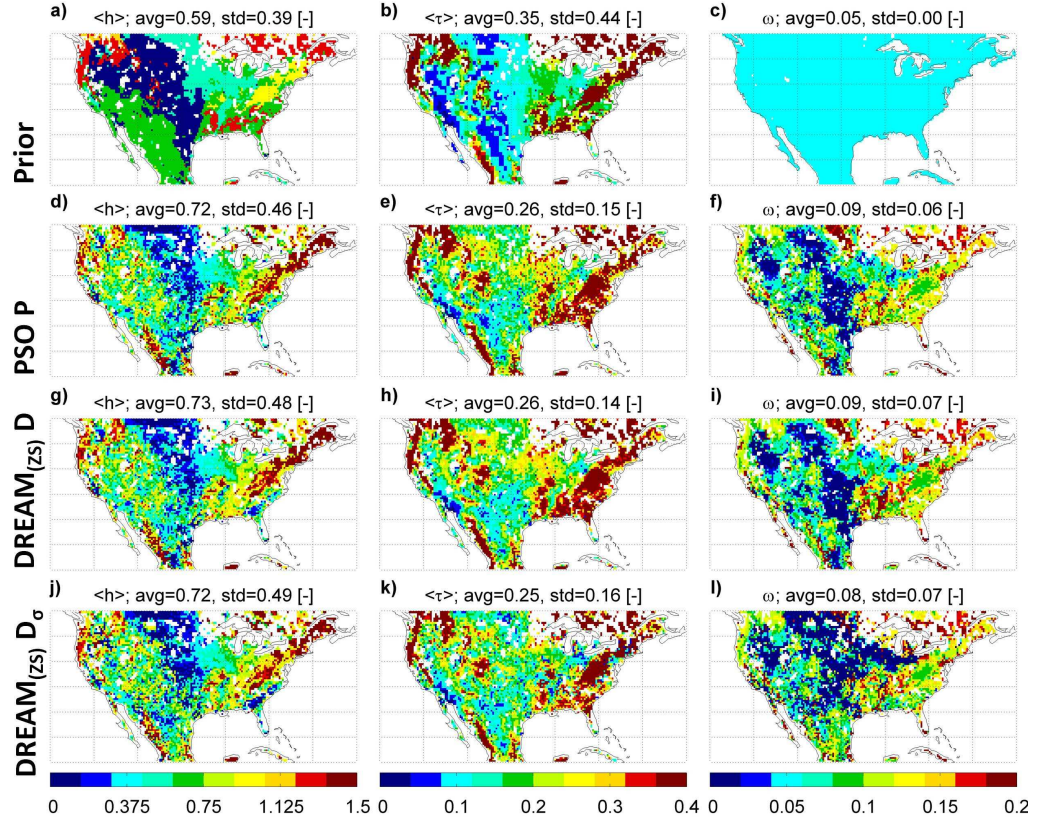


Figure 3: Parameter values for (left) $\langle h \rangle$, (middle) $\langle \tau \rangle$, and (right) ω , for the (top row) prior distribution and (second row) case P, (third row) case D and (fourth row) case D _{σ} .

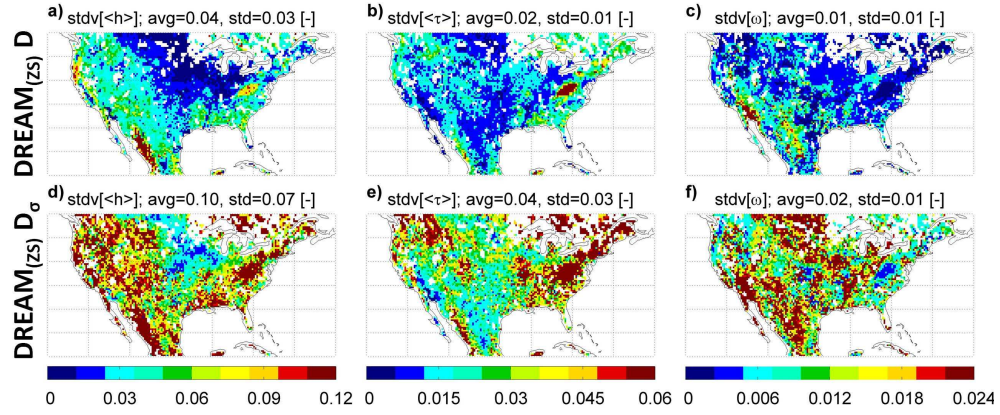


Figure 4: Uncertainty in parameter estimates for (left) $\langle h \rangle$, (middle) $\langle \tau \rangle$, and (right) ω , obtained with DREAM_(ZS) (top row) case D and (bottom row) case D_σ. Case D underestimates the posterior parameter uncertainty (see text).

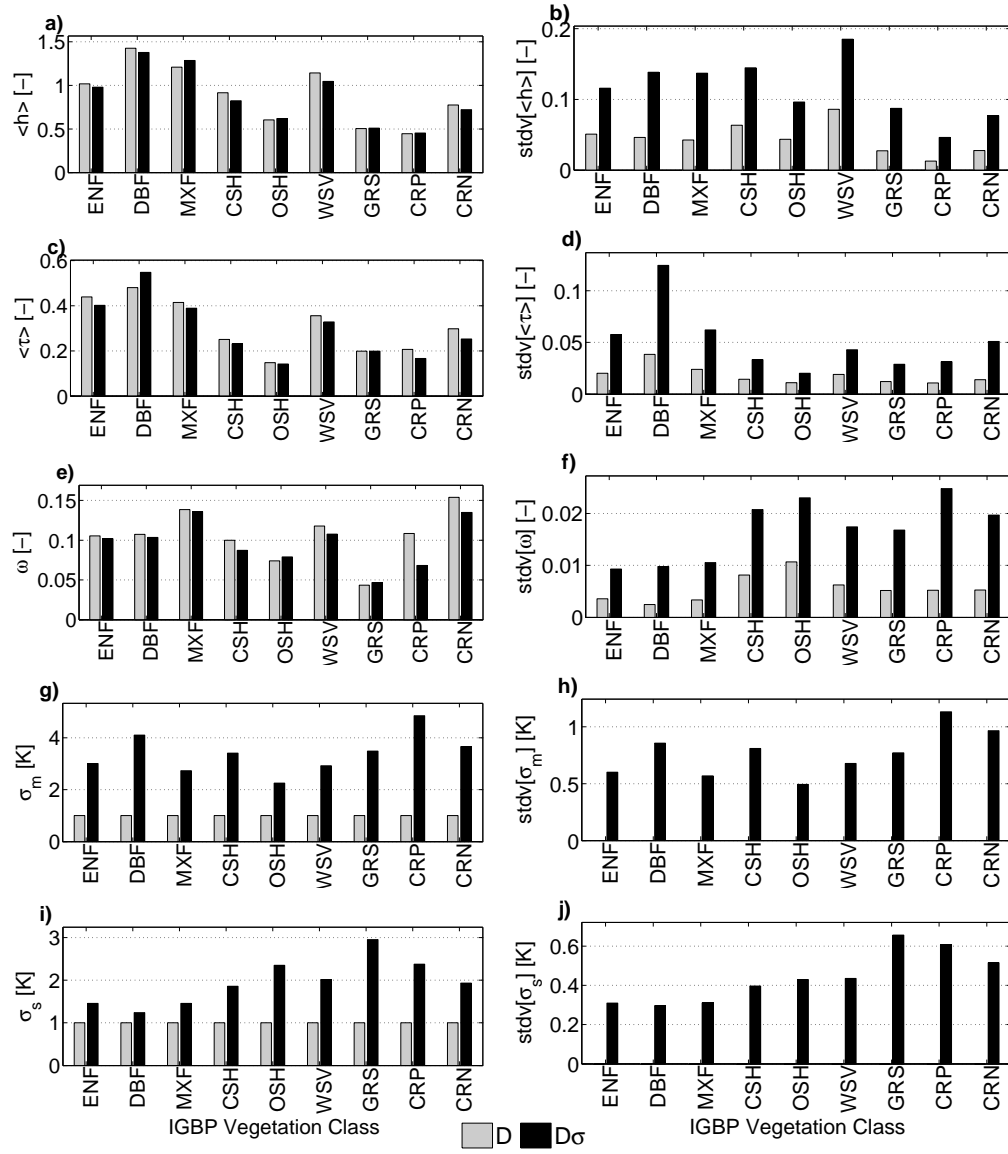


Figure 5: (Left) MAP parameter values and (right) uncertainties aggregated per vegetation class for DREAM_(ZS) cases D and D_σ. Each row represents a different parameter: (a,b) $\langle h \rangle$, (c,d) $\langle \tau \rangle$, (e,f) ω , (g,h) σ_m , (i,j) σ_s .

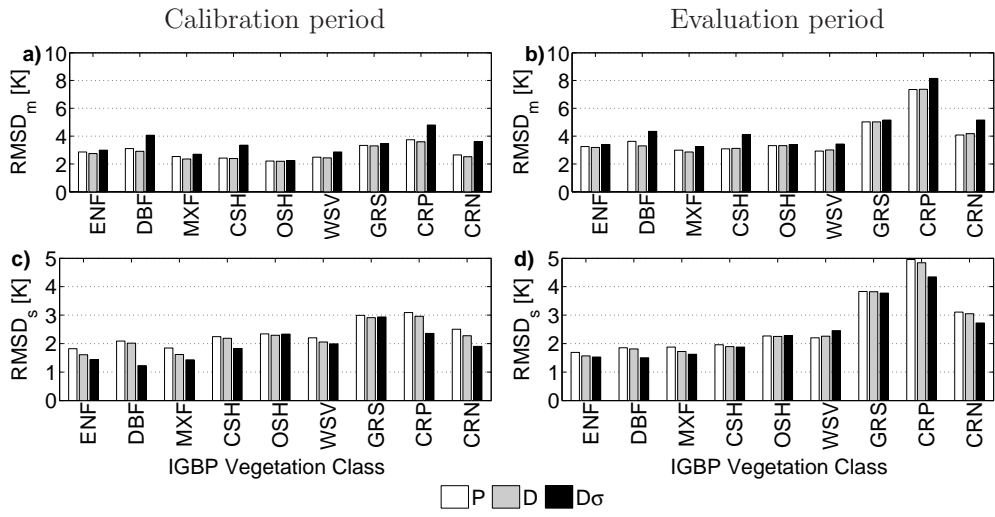


Figure 6: RMSD in long-term Tb (a,b) average and (c,d) standard deviation during the (a,c) calibration period (1 July 2011 - 1 July 2012) and (b,d) evaluation period (1 July 2010 - 1 July 2011), using the MAP parameter values derived from PSO (case P) and DREAM_(ZS) (cases D and D_σ).

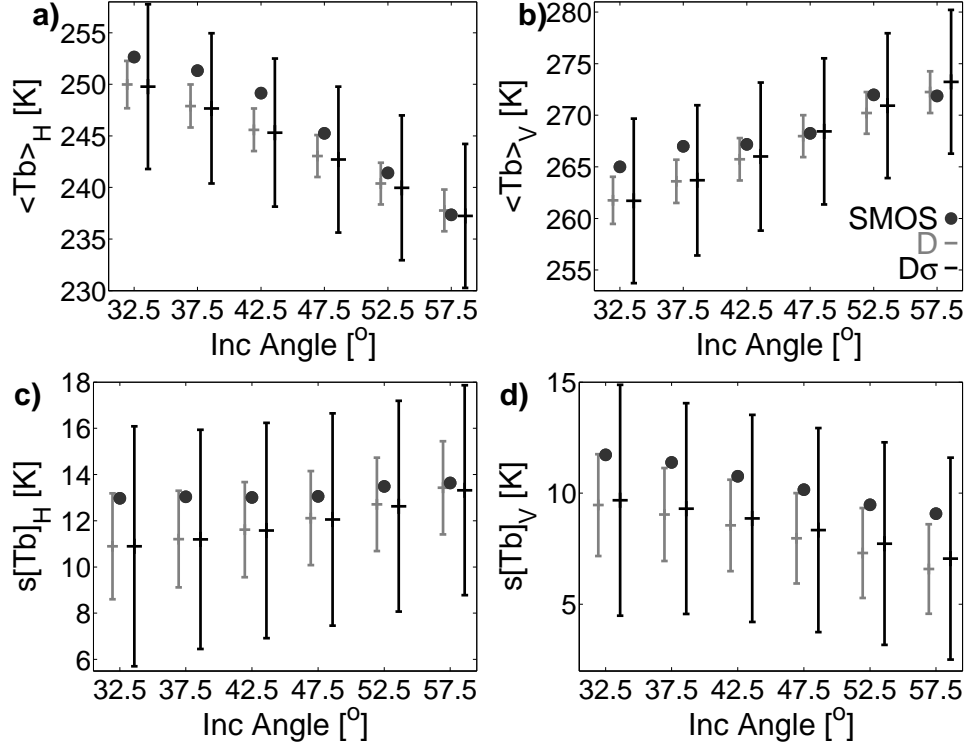


Figure 7: (a,b) Long-term average and (c,d) standard deviation, for (a,c) H- and (b,d) V-polarized Tb from (dots) SMOS observations and (horizontal dashes) simulations averaged over the study domain, during the calibration period (1 July 2011 - 1 July 2012) and only including morning overpasses. The Tb simulations use MAP parameter estimates derived with DREAM_(ZS) (gray) case D and (black) case D σ , and are shown by a central horizontal dash. The error bars indicate (a,b) $\pm 2\sigma_{i,m}$ and (c,d) $\pm 2\sigma_{i,s}$, i.e twice the lumped standard deviations of the residual simulation and observation errors, and are drawn around the simulated Tb for illustration. For clarity, symbols are slightly offset from the nominal incidence angle.

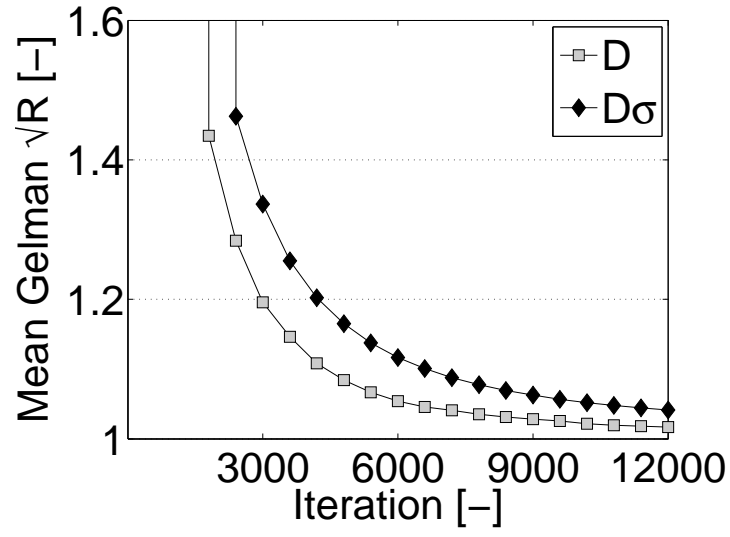


Figure 8: Gelman-Rubin convergence diagnostic \sqrt{R} for the two DREAM_(ZS) MCMC simulation cases. The metric is averaged over all calibrated parameters, and across the study domain.